

A 2-D extension of the sampling algorithm for sparse fourier representations

Kazutaka Nishimura,
Yoshinori Takei, Toshinori Yoshikawa,
Dept. Electrical Engineering
Nagaoka University of Techlogy
1603-1 Kamitomikamachi
Nagaoka-shi, Niigata 940-2188

and

Xi Zhang

Dept. Information and Communication Engineering
The University of Electro-Communications
1-5-1 Chofugaoka
Chofu-shi, Tokyo 182-8585

Abstract—Recently, a very efficient sampling algorithm for finding a B -term fourier representation of given 1-D discrete signal is presented [Gilbert, Guha, Indyk, Muthukrishnan, Strauss; STOC02]. In this paper, we present a modified version of the algorithm, which can be applied to 2-D signals. As in the original algorithm, dependence of the running time on the signal length is polylogarithmic.

I. INTRODUCTION

Let

$$\mathbf{x} = (x[0], x[1], \dots, x[n], \dots, x[N-1]) \quad (1)$$

be a complex-valued discrete time signal of length N . The discrete fourier transform (DFT), or the spectrum, of \mathbf{x} ,

$$\hat{\mathbf{x}} = (\hat{x}[0], \hat{x}[1], \dots, \hat{x}[k], \dots, \hat{x}[N-1]) \quad (2)$$

where

$$\hat{x}[k] := \frac{1}{\sqrt{N}} \sum_{n=0}^{N-1} x[n] e^{-j2\pi nk/N}, \quad (3)$$

is a fundamental tool for analyzing or representing features of the signal \mathbf{x} . The inverse DFT (IDFT):

$$x[n] = \frac{1}{\sqrt{N}} \sum_{k=0}^{N-1} \hat{x}[k] e^{j2\pi nk/N}, \quad (4)$$

recovers the signal \mathbf{x} completely from the spectrum $\hat{\mathbf{x}}$, and these transforms preserve the ℓ^2 norm:

$$\|\mathbf{x}\|_2 = \|\hat{\mathbf{x}}\|_2, \quad (5)$$

where

$$\|\mathbf{x}\|_2 := \sqrt{\sum_{n=0}^{N-1} |x[n]|^2}. \quad (6)$$

Let $B < N$ be a positive integer and suppose that the frequency list $K(B) = (k_1, \dots, k_B)$ is taken so that $\hat{x}[k_1], \dots, \hat{x}[k_B]$ are the B -maxima among the spectrum $\hat{\mathbf{x}}$, in their absolute values. Then the signal $\mathbf{r}_{\text{opt}(B)}$ defined by

$$r_{\text{opt}(B)}[n] = \frac{1}{\sqrt{N}} \sum_{k \in K(B)} \hat{x}[k] e^{j2\pi nk/N} \quad (7)$$

is the best possible ℓ^2 approximation of the signal \mathbf{x} by a B -term sum of pure tones $\psi_k[n] := e^{j2\pi nk/N}/\sqrt{N}$. We shall refer $\mathbf{r}_{\text{opt}(B)}$ as the optimal B -term (fourier) representation of \mathbf{x} .

In those applications that the optimal B -term representation permit a good approximation to the original signal with $B \ll N$, the B -term representation, or equivalently, the lists $K(B)$ of frequencies and $\hat{x}[k_1], \dots, \hat{x}[k_B]$ of coefficients, would be a nice “digest” of the original signal, in the views of conciseness and accuracy. The computation is obviously feasible, by a combination of the fast fourier transform (FFT) algorithm and some sorting algorithm. The running time has a lower bound $\Omega(N \log N)$ (here we regard the required precision as a constant), which sounds not so big. However, when we want to make B -term digest for each of \mathbf{x} ’s, in a huge database, each having a really long length N , the cost $N \log N$ by the number of \mathbf{x} ’s will be so expensive. The formula (3) shows that every fourier coefficient $\hat{x}[k]$ depends on entire signal \mathbf{x} , and it would be natural that one think it were impossible to reduce the computational complexity even with the sparseness condition $B \ll N$ (say $B = 10$ independent of N .)

Gilbert, Guha, Indyk, Muthukrishnan, and Strauss [1] have shown a somewhat surprising result that a *near-optimal* B -term representation of any signal can be obtained *with high probability*, within a time whose dependence on N is poly-logarithmic. Let ι be a small number such that, if $\|\mathbf{x} - \mathbf{y}\|_2^2 \leq \iota$ then we consider that two signals \mathbf{x} and \mathbf{y} are identical. By a normalization, we assume that $\iota = 1$ and M is a crude upper bound on $\|\mathbf{x}\|_2$ after the normalization.

Theorem 1 ([1]). *There exists an algorithm, on input $B < N$, $\epsilon > 0$, and \mathbf{x} , with cost $(B \log(N) \log(M)/\epsilon)^{O(1)}$, with high probability, output a B -term representation \mathbf{r} for \mathbf{x} which satisfies $\|\mathbf{x} - \mathbf{r}\|_2 \leq (1 + \epsilon) \|\mathbf{x} - \mathbf{r}_{\text{opt}(B)}\|_2$, where $\mathbf{r}_{\text{opt}(B)}$ is the optimal B -term representation for \mathbf{x} . The algorithm accesses only $(B \log(N) \log(M)/\epsilon)^{O(1)}$ samples $\{x[n] : n \in T\}$ of \mathbf{x} , where the random sample set T is chosen independently of \mathbf{x} . (The success probability may be a constant arbitrarily close to 1 and the cost for it is implied in “ $O(1)$ ”.)*

Thus, when B is sufficiently smaller than N , the algorithm, which we shall refer as the GGIMS algorithm in the sequel, enables very efficient computation of B -term representations.

Now, suppose that \mathbf{x} is a signal defined over the 2-dimensional domain $\{0, \dots, N_1 - 1\} \times \{0, \dots, N_2 - 1\}$:

$$\mathbf{x} = (x[n_1, n_2] : 0 \leq n_1 < N_1, 0 \leq n_2 < N_2) \quad (8)$$

and let

$$\hat{\mathbf{x}} = (\hat{x}[k_1, k_2] : 0 \leq k_1 < N_1, 0 \leq k_2 < N_2) \quad (9)$$

be its 2-D DFT, where

$$\hat{x}[k_1, k_2] := \sum_{n_1, n_2} x[n_1, n_2] \psi_{k_1, k_2}[-n_1, -n_2], \quad (10)$$

and

$$\psi_{k_1, k_2}[n_1, n_2] := \frac{e^{j2\pi(\frac{k_1 n_1}{N_1} + \frac{k_2 n_2}{N_2})}}{\sqrt{N_1 N_2}}. \quad (11)$$

The optimal B -term representation for \mathbf{x} , i.e. the lists $K(B) = ((k_{1,1}, k_{2,1}), \dots, (k_{1,B}, k_{2,B}))$ of frequencies and the corresponding coefficients $\hat{x}[k_{1,1}, k_{2,1}], \dots, \hat{x}[k_{1,B}, k_{2,B}]$ where the latter are B -maxima of \hat{x} in absolute value, will be useful in this case also. The goal of this paper is to extend the GGIMS algorithm to an algorithm that efficiently compute near-optimal B -term representations for signals defined over a 2-dimensional domain.

The rest of paper is organized as follows: Section II. is a brief description of the GGIMS algorithm [1] for 1-dimensional signals, where we shall discuss what parts of the algorithm are dimension-dependent. In Section III., we modify parts of the GGIMS algorithm so that it can be applied for 2-dimensional signals. Section IV. concludes the paper.

In the sequel, the time or frequency domain $\{0, 1, \dots, N-1\}$ is identified with $\mathbf{Z}/N\mathbf{Z}$, the ring of integers modulo N . For simplicity, we assume that N, N_1, N_2 are odd prime numbers. $\mathbb{E}[X]$ means the expectation of a random variable X .

II. THE GGIMS ALGORITHM

The GGIMS algorithm composed of 3 parts; Identification, Estimation and Iteration. Given the input signal \mathbf{x} , the algorithm start with the current list S set to null, and the corresponding at-most- B -term representation \mathbf{r} is set to 0.

A. Identification

Let $\eta := \epsilon/(2B)$. The purpose of this part is to output a list Λ of frequencies, whose length is $2m+1 = O(1/\eta)$, which contains (with high probability) all frequencies k satisfying $|\hat{x}[k]|^2 \geq \eta \|\mathbf{x}\|_2^2$. In this part there exist two procedures; Isolation and Group Testing. The latter procedure call a subroutine that (roughly) estimate the norm of given signal by sampling.

1) *Isolation*: This procedure constructs $2m+1$ signals $\mathbf{f}_0, \dots, \mathbf{f}_{2m}$ that satisfy (with high probability)

1. For each k' such that $|\hat{x}[k']|^2 \geq \eta \|\mathbf{x}\|_2^2$, there exists $i(0 \leq i \leq 2m+1)$ such that $\hat{f}_i[k'] \geq 0.98 \|\mathbf{f}_i\|_2^2$.
2. Each \mathbf{f}_i can be sampled by sampling non-adaptively from \mathbf{x} in $O(m)$ places.

Here we note that the word “constructs” does not mean actual output of the entire signal points $(f_i[0], f_i[1], \dots, f_i[N-1])$.

The procedure only sets up parameters which enable later procedures to sample arbitrary signal points of \hat{f}_i . To show the construction we define the operator $R_{\theta, \sigma}$ for $\sigma, \theta \in \mathbf{Z}/N\mathbf{Z}$ through

$$(R_{\theta, \sigma} f)[n] = e^{j2\pi \theta n/N} f[\sigma n]. \quad (12)$$

Then it holds that

$$(\widehat{R_{\theta, \sigma} f})[\sigma k + \theta] = \hat{f}[k]. \quad (13)$$

Also we define the Fejér kernel of length $2m+1$ as

$$H_m[n] := \begin{cases} \frac{\sqrt{N}}{2m+1} & n \in [-m, m] \\ 0 & \text{otherwise.} \end{cases} \quad (14)$$

It holds that

$$\hat{H}_m[k] = \begin{cases} \frac{\sin(\pi(2m+1)/N)}{(2m+1) \sin(\pi k/N)} & (k \neq 0) \\ 1 & (k = 0) \end{cases} \quad (15)$$

and in particular $\hat{H}_m[k] \geq 2/\pi$ for $k \in [-N/(2(2m+1)), N/(2(2m+1))]$. We regard the last interval as “pass region” of \mathbf{H}_m

Now, the signals $\mathbf{f}_0, \dots, \mathbf{f}_{2m}$ are defined as follows: first, pick θ and σ at random in $\mathbf{Z}/N\mathbf{Z}$, with σ invertible. Then put, for each $0 \leq i \leq 2m$,

$$\mathbf{f}_i := (e^{j2\pi i n/(2m+1)} \mathbf{H}_m) * (R_{\theta, \sigma} \mathbf{x}), \quad (16)$$

where $*$ denotes the convolution. In the frequency domain, the spectrum \hat{x} is first permuted by $k \mapsto \sigma k + \theta$, then filtered by “ $iN/(2m+1)$ -shifted version of” $\hat{\mathbf{H}}_m$ to form $\hat{\mathbf{f}}_i$.

Using the fact that $k \mapsto \sigma k + \theta$ is a pair-wise independent permutation (i.e., for all $k_1 \neq k_2$ and $k_3 \neq k_4$, $k_1 \mapsto k_3$ and $k_2 \mapsto k_4$ with probability $1/(N(N-1))$), it can be shown that the following holds with high probability: for each k' such that $|\hat{x}[k']|^2 \geq \eta \|\mathbf{x}\|_2^2$, there exists $i(0 \leq i \leq 2m+1)$ such that $\hat{f}_i[k'] \geq 0.98 \|\mathbf{f}_i\|_2^2$. We note that sampling one point from \mathbf{f}_i can be done by sampling $2m+1$ points from \mathbf{x} , by construction.

2) *Group Testing*: For each $\mathbf{f} \in \{\mathbf{f}_0, \dots, \mathbf{f}_{2m}\}$, the procedure determines the frequency k' satisfying $\hat{f}[k'] \geq 0.98 \|\mathbf{f}\|_2^2$, if such k' exists. Then gather such k' from \mathbf{f}_i 's to form the frequency list Λ . By the construction of \mathbf{f}_i 's, Λ catches all such frequencies that $|\hat{x}[k']|^2 \geq \eta \|\mathbf{x}\|_2^2$.

The procedure utilizes 16 filters

$$\hat{G}_\ell[k] = \frac{1}{2} (1 + \cos(\frac{2\pi k}{N} - \frac{2\pi \ell}{16})) \quad (\ell = 0, \dots, 15), \quad (17)$$

each is of length 3. If we define the pass region of $\hat{\mathbf{G}}_\ell$ by

$$\text{pass}_\ell := \{k : |2\pi k/N - 2\pi \ell/16| \leq 2\pi/32\}, \quad (18)$$

then $\hat{G}_\ell[k] \geq 0.99$ for $k \in \text{pass}_\ell$ and these pass regions tile the entire frequency domain. The procedure also uses a subroutine, which will be explained in the next subsection, satisfying the

```

 $c \leftarrow 2M/\beta$ 
Do
{
    Pick  $r$  random independent samples  $n_i \in \mathbf{Z}/N\mathbf{Z}$ 
     $X \leftarrow \frac{1}{r} \sum_{i=1}^r N|(K_c f)[n_i]|^2$ 
     $c \leftarrow c/(1 + \beta)$ 
} while  $X < \beta^2 c^2$  and  $c \geq 1$ 
Output  $X/(1 + \beta)$ 

```

Fig. 1. Estimating Norm

following: The subroutine makes $O(\log \log M)$ samples, runs in time polynomial in $\log(M)$, on input \mathbf{f} (of length N) with biggest fourier coefficient $\hat{f}[k]$, it returns, with high probability, a random output X such that:

1. $X \leq \|\mathbf{f}\|_2^2, \forall \mathbf{f}$.
2. If $|\hat{f}[k]|^2 \geq 0.95\|\mathbf{f}\|_2^2$ then $X \geq 0.5\|\mathbf{f}\|_2^2$.

Now assume that we are seeking unknown k' which have at least 98% energy of \mathbf{f} . Without loss of generality, we assume that $k' \in \text{pass}_0$. For each $\ell = 0, 1, \dots, 15$, We estimate the norm $\|\mathbf{f} * G_\ell\|$ using the subroutine. For $\ell = 0$, it can be shown that $|\mathbf{f} * \hat{G}_0|^2 \geq 0.95\|\mathbf{f} * G_0\|_2^2$ by combining the facts $k' \in \text{pass}_0$, $\hat{G}_0[k'] \geq 0.99$ and $\hat{f}[k'] \geq 0.98\|\mathbf{f}\|_2^2$. Then the second property of the subroutine, the returned value X satisfies $X \geq 0.5\|\mathbf{f} * \hat{G}_0\|_2^2$. It also can be shown that $0.5\|\mathbf{f} * \hat{G}_0\|_2^2 \geq 0.48\|\mathbf{f}\|_2^2$. On the other hand, for $\ell = 4$, it can be shown that the returned value X of the subroutine satisfies $X \leq \|\mathbf{f} * \hat{G}_4\|_2^2 \leq 0.38\|\mathbf{f}\|_2^2$, from the facts that k' is far from pass_4 and that the contributions to $\|\mathbf{f} * \hat{G}_4\|_2^2$ from frequencies other than k' is small (at most $0.02\|\mathbf{f}\|_2^2$). For $\ell = 5, 6, \dots, 12$, we obtain $X \leq 0.38\|\mathbf{f}\|_2^2$ in the same way. So, we can eliminate the possibilities that $k' \in \text{pass}_\ell$ for $\ell = 4, \dots, 12$. In this manner, we can always eliminate 9/16 possibilities. The remaining region forms a cyclic interval of length at most $7N/16$. By applying the operator $R_{0,2}$, we can dilate the remaining frequency region by two. Then applying 16 filters, we can halve the possibilities again. Repeating that $O(\log(N))$ times, we learn k' .

3) *Estimating Norm*: The subroutine for estimating norm is shown in Fig. 1, where $\beta > 0$ is a small constant (independent of N and M), and the number r of samples is $O(\log \log M)$, and in fact we can reuse the same samples in the loop. $K_c f$ stands means “ f clipped at the ceiling c ”, i.e.,

$$(K_c f)[n] := \begin{cases} f[n] & (|f[n]| \leq c/\sqrt{N}) \\ 0 & \text{otherwise.} \end{cases} \quad (19)$$

Intuition behind the algorithm is as follows: Let $Y = \frac{1}{r} \sum_{i=1}^r N|f[n_i]|^2$. Then $E[Y] = \|\mathbf{f}\|_2^2$ and it would be a good estimate if the variance of Y were small. The variance comes from spikes, so we use the clipped version X instead. It is satisfied that $0 \leq E[X] \leq E[Y] = \|\mathbf{f}\|_2^2$ and X has small variance.

If 95% of \mathbf{f} is concentrated in a pure frequency, then the energy of spikes are small, so $E[X] \approx E[Y] = \|\mathbf{f}\|_2^2$, and since variance of X is small, X is a good estimation. The algorithm sets a ceiling c , above which all values of the function are clipped as if they were spikes. The algorithm gradually lowers the clipping ceiling c until the energy estimate is consistent with the ceiling value.

B. Estimation

This part estimates $\hat{x}[k]$, for each $k \in \Lambda$, by a sampling algorithm.

The sampling algorithm used here takes inputs (\mathbf{a}, k, μ) , where \mathbf{a} is a signal of length N , k a frequency, μ an accuracy parameter. Then the algorithm runs with cost $O(\log(M)/\mu)$, making $O(1/\mu)$ samplings from \mathbf{a} , and returns the estimation $\hat{a}[k]$, which satisfies with high probability that $|\hat{a}[k] - a[k]|^2 \leq \epsilon\|\mathbf{a}\|_2^2$. The estimation $\hat{a}[k]$ is computed as follows: let $t = O(1/\mu)$ and pick t random positions n_1, \dots, n_t . Put $\mathbf{a}'' = N/t \sum_{h=1}^t \Delta_{n_h}$, where Δ_n is the delta function. Then compute $X = \mathbf{a}''[k]$ and return it as the estimate. It can be shown that $E[X] = a[k]$ and $E[|X - a[k]|^2] \leq O(\epsilon\|\mathbf{a}\|_2^2)$.

For the estimation of $\hat{x}[k]$, an accuracy of $\mu = 1/\text{poly}(B \log(N) \log(M/\delta)/\epsilon)$ (where $\delta = \|\mathbf{x} - \mathbf{r}_{\text{opt}(B)}\|_2$) is required.

C. Iteration

Now we have the estimated fourier coefficients $\hat{x}[k]$ for $k \in \Lambda$. Let $\tilde{x}[k']$ be the maximum (in absolute value) among them. Add k' to the current coefficients list S , and update the current representation \mathbf{r} by $\mathbf{r} \leftarrow \mathbf{r} + \tilde{x}[k']\psi_{k'}$. Then set $\mathbf{x} \leftarrow \mathbf{x} - \tilde{x}[k']\psi_{k'}$ and A.. However, we should note that the update of \mathbf{x} just before goto is in fact impossible, since we can access \mathbf{x} only via sampling. Actually, after 2nd iteration, we sample from $\mathbf{x} - \mathbf{r}$ whenever we need sample from \mathbf{x} , using the maintained representation \mathbf{r} .

The algorithm may or may not halt just after that the representation \mathbf{r} has grown to B -terms. Here we do not mention the accurate halting condition, however, it is guaranteed to halt within $\text{poly}(B \log(N) \log(M/\delta)/\epsilon)$ iterations.

III. MODIFICATIONS FOR 2-D SIGNALS

The strategy of GGIMS algorithm is not dimension-dependent, However, two key steps in the Identification part needs a modification for 2-D signals. Here we describe modifications for Isolation step and Group Testing step. In the sequel, signals are defined over the 2-dimensional domain $\mathbf{Z}/N_1\mathbf{Z} \times \mathbf{Z}/N_2\mathbf{Z}$.

A. Identification (2-D)

1) *Isolation (2-D)*: Let $\eta := \epsilon/(2B)$. We will construct $(2m_1 + 1) \times (2m_2 + 1)$ signals $f_{i_1, i_2}[n_1, n_2](i_1 =$

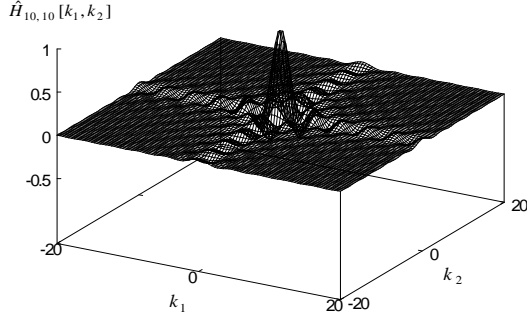


Fig. 2. $\hat{H}_{10,10}[k_1, k_2]$ ($N_1 = N_2 = 40$)

$0, 1, \dots, 2m_1, i_2 = 0, 1, \dots, 2m_2$), where $(2m_1 + 1) \times (2m_2 + 1) = O(1/\eta)$.

We define the operator $R_{\theta_1, \sigma_1, \theta_2, \sigma_2}$ through

$$R_{\theta_1, \sigma_1, \theta_2, \sigma_2} x[n_1, n_2] = e^{j2\pi(\theta_1 n_1/N_1 + \theta_2 n_2/N_2)} x[\sigma_1 n_1, \sigma_2 n_2] \quad (20)$$

and 2-D Fejér kernel by

$$H_{m_1, m_2}[n_1, n_2] = \begin{cases} \frac{\sqrt{N_1 N_2}}{(2m_1 + 1)(2m_2 + 1)} & (n_1, n_2) \in [-m_1, m_1] \times [-m_2, m_2] \\ 0 & (\text{otherwise}) \end{cases} \quad (21)$$

See Fig.2 for the spectrum.

Then the signals $f_{i_1, i_2}[n_1, n_2]$ ($i_1 = 0, 1, \dots, 2m_1, i_2 = 0, 1, \dots, 2m_2$) are constructed as follows: pick $\sigma_1, \theta_1 \in \mathbf{Z}/N_1\mathbf{Z}$ at random, with σ_1 invertible. Independently, pick $\sigma_2, \theta_2 \in \mathbf{Z}/N_2\mathbf{Z}$ at random, with σ_2 invertible. Then set

$$f_{i_1, i_2}[n_1, n_2] := (e^{j2\pi(i_1 n_1/(2m_1 + 1) + i_2 n_2/(2m_2 + 1))} H_{m_1, m_2}[n_1, n_2]) * R_{\theta_1, \sigma_1, \theta_2, \sigma_2} x[n_1, n_2]. \quad (22)$$

It can be shown that they have the following (expected) properties:

1. For each k'_1, k'_2 such that $|\hat{x}[k'_1, k'_2]|^2 \geq \eta \|\mathbf{x}\|_2^2$, there exists (i_1, i_2) ($0 \leq i_1 \leq 2m_1 + 1, 0 \leq i_2 \leq 2m_2 + 1$) such that $\hat{f}_{i_1, i_2}[k'_1, k'_2] \leq 0.98 \|\mathbf{f}_{i_1, i_2}\|_2^2$.
2. Each \mathbf{f}_{i_1, i_2} can be sampled by sampling non-adaptively from \mathbf{x} in $O(m_1 m_2)$ places.

2) *Group Testing (2-D)*: In 2-D case, we rely on the following 256 filters (see Fig. 3):

$$\hat{G}_{\ell_1, \ell_2}[k_1, k_2] = \frac{1}{4} \left\{ 1 + \cos\left(\frac{2\pi k_1}{N_1} - \frac{2\pi \ell_1}{16}\right) \right\} \cdot \left\{ 1 + \cos\left(\frac{2\pi k_2}{N_2} - \frac{2\pi \ell_2}{16}\right) \right\} \quad (0 \leq \ell_1 < 16, 0 \leq \ell_2 < 16) \quad (23)$$

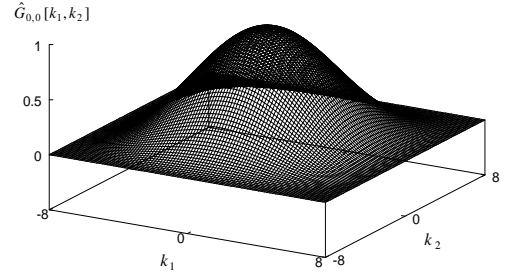


Fig. 3. The filter $\hat{G}_{0,0}[k_1, k_2]$ ($N_1 = N_2 = 16$)

In the time domain, $\mathbf{G}_{\ell_1, \ell_2}$ has a support of 3×3 square.

With these filters, we can eliminate $1 - (7/16)^2 = 207/256$ (instead of $1 - 7/16$ in 1-D case) of possibilities in one step of Group Testing.

B. Complexity of 2-D algorithm

By the above modifications of the algorithm, we obtain a natural 2-D extension of Theorem 1:

Theorem 2. *There exists an algorithm, on input $B < N_1 N_2$, $\epsilon > 0$, and a signal \mathbf{x} over the 2-dimensional domain $\{0, \dots, N_1 - 1\} \times \{0, \dots, N_2 - 1\}$, with cost $(B \log(N_1) \log(N_2) \log(M)/\epsilon)^{O(1)}$, with high probability, output a B -term representation \mathbf{r} for \mathbf{x} which satisfies $\|\mathbf{x} - \mathbf{r}\|_2 \leq (1 + \epsilon) \|\mathbf{x} - \mathbf{r}_{\text{opt}(B)}\|_2$, where $\mathbf{r}_{\text{opt}(B)}$ is the optimal B -term representation for \mathbf{x} . The algorithm accesses only $(B \log(N_1) \log(N_2) \log(M)/\epsilon)^{O(1)}$ samples $\{x[n_1, n_2] : (n_1, n_2) \in T\}$ of \mathbf{x} , where the random sample set T is chosen independently of \mathbf{x} . (The success probability may be a constant arbitrarily close to 1 and the cost for it is implied in “ $O(1)$ ”.)*

IV. CONCLUSION

In this paper, we have presented an efficient algorithm that finds a near-optimal B -term fourier representation of 2-D signals, by modifying a part of GGIMS algorithm [1].

The running time of the algorithm depends on signal amount only polylogarithmically, so the algorithm will be useful for digesting vast 2-D signals.

The algorithm has many hidden constants that would affect practical use. So more fine analysis of the algorithm is still needed.

REFERENCES

- [1] A. C. Gilbert, S. Guha, P. Indyk, S. Muthukrishnan, and M. Strauss, “Near-Optimal Sparse Fourier Representations via Sampling,” *In Proceedings of the 34th Annual ACM Symposium on Theory of Computing (STOC)*, pp.152-161, May 2002.